

1 **Seasonal predictability of daily rainfall characteristics in central-**
2 **northern Chile for dry-land management**

3
4
5
6 Koen Verbist

7 International Centre for Eremology, Dep. Soil Management, Ghent University, Ghent,
8 Belgium

9 Water Centre for Arid Zones in Latin America and the Caribbean (CAZALAC), La Serena,
10 Chile

11
12 Andrew W. Robertson

13 International Research Institute for Climate and Society, Palisades, New York, USA

14
15 Wim M. Cornelis

16 International Centre for Eremology, Dep. Soil Management, Ghent University, Ghent,
17 Belgium

18
19 Donald Gabriels

20 International Centre for Eremology, Dep. Soil Management, Ghent University, Ghent,
21 Belgium

22
23
24 Note: The authors indicate the use of nonstandard references in the introduction of this
25 manuscript that deal with the unpublished drought information of the region under study.

26
27
28
29
30 Revised version: February 3, 2010

31
32 *Journal of Applied Meteorology and Climatology, sub judice*

33
34
35
36
37
38
39
40
41
42

Corresponding author address: Koen Verbist, Dep. Soil Management, Coupure Links 653, 9000 Ghent
43 Belgium.

44 E-mail: Koen.Verbist@UGent.be

Abstract

45

46 The seasonal predictability of daily winter rainfall characteristics relevant to dry-land
47 management was investigated in the Coquimbo Region of central-northern Chile, with focus
48 on the seasonal rainfall total, daily rainfall frequency, and mean daily rainfall intensity on wet
49 days at the station scale. Three approaches of increasing complexity were tested. First, an
50 index of the simultaneous El Niño-Southern Oscillation (ENSO) was regressed onto May-
51 August (MJJA) observed precipitation; this explained 32% of station-averaged rainfall-
52 amount variability, but performed poorly in a forecasting setting. The second approach used
53 retrospective seasonal forecasts made with three general circulation models (GCMs) to
54 produce downscaled seasonal rainfall statistics by means of canonical correlation analysis
55 (CCA). In the third approach, a non-homogeneous Hidden Markov Model (nHMM) driven
56 by the GCM's seasonal forecasts was used to model stochastic daily rainfall sequences. While
57 the CCA is used as a downscaling method for the seasonal rainfall characteristics themselves,
58 the nHMM has the ability to simulate a large ensemble of daily rainfall sequences at each
59 station from which the rainfall statistics were calculated. Similar cross-validated skill
60 estimates were obtained using both the CCA and nHMM, with the highest correlations with
61 observations found for seasonal rainfall amount and rainfall frequency (up to 0.9 at individual
62 stations). These findings were interpreted using analyses of observed rainfall spatial
63 coherence, and by means of synoptic rainfall states derived from the HMM. The downscaled
64 hindcasts were then tailored to meteorological drought prediction, using the Standardized
65 Precipitation Index (SPI) based on seasonal values, and the frequency of substantial rainfall
66 days (>15mm; FREQ15) and the daily Accumulated Precipitation Deficit. Deterministic
67 hindcasts of SPI showed high hit rates, with high Ranked Probability Skill Score for
68 probabilistic hindcasts of FREQ15 obtained via the nHMM.

69 **1. Introduction**

70 Climate variability can have serious social impacts in semi-arid regions, especially for
71 farmers who depend on rain-fed agriculture and on livestock production based on natural
72 vegetation. In the Coquimbo Region in central-northern Chile, where rainfall amounts often
73 drop under the limit for crop growth, a lack of rainfall results in a crisis situation for society.
74 Over 2.6 million US Dollars was spent during 2007 to support affected families and farmers
75 in the Coquimbo Region, to repair damage, to recover degraded soils and to increase irrigation
76 programs (Chilean Ministry of Agriculture 2008, personal communication). Although these
77 measures reduced the negative effects of the 2007 drought, they did not address all affected
78 families due to budget limitations, nor did they increase preparedness and resilience to future
79 droughts. Of the 16 307 rural families in Chile seeking monetary aid to overcome the negative
80 aspects of the 2007 drought, more than 75% indicated suffering a lack of sufficient water for
81 irrigation, for domestic use, and they observed harvest losses for the crops grown for their
82 own consumption (Fondo de Solidaridad e Inversión Social (FOSIS) 2008, personal
83 communication). A typical problem here is the lack of preparedness prior to these natural
84 events, making any governmental action afterwards less cost effective. Despite the need, the
85 current Drought Alleviation Plan formulated by the Chilean Government for the region
86 (FOSIS 2008) does not include strategies for drought early warning, and the feasibility of
87 such a system has yet to be demonstrated.

88 The El Niño Southern Oscillation (ENSO) is known to have a strong impact on winter rainfall
89 over central-northern Chile, with positive rainfall anomalies during El Niño events, and below
90 normal rainfall mostly associated with La Niña conditions (Aceituno 1988; Aceituno et al.
91 2009; Falvey and Garreaud 2007; Montecinos and Aceituno 2003; Pittock 1980; Quinn and

92 Neal 1983; Rubin 1955; Rutllant and Fuenzalida 1991; Garreaud et al. 2009). However, the
93 associated seasonal predictability and forecast skill levels from current dynamical seasonal
94 prediction models (e.g. Goddard et al. 2003) have not yet been assessed in detail for the
95 statistics of local daily weather that are likely to be most pertinent to meteorological drought.

96 In this paper we document the characteristics of daily winter rainfall from station
97 observations over the Coquimbo Region, and assess their seasonal predictability from three
98 current seasonal prediction general circulation models (GCMs), together with statistical
99 techniques to “downscale” and tailor the output from these relatively coarse resolution models
100 to the station scale. While GCMs typically misrepresent the characteristics of local daily
101 rainfall, statistical downscaling can often correct such biases and provide probabilistic rainfall
102 simulations that are well calibrated against local station data (Hughes and Guttorp 1994;
103 Robertson et al. 2009). Our analysis of the station rainfall data begins with a decomposition of
104 seasonal rainfall amounts into rainfall frequency and the mean rainfall amount falling on wet
105 days, i.e. the rainfall intensity. The correlation between rainfall data from stations separated
106 by increasing distances, i.e. the spatial “coherence,” for each of the seasonal anomaly types
107 (rainfall amount, intensity and frequency) across the region is then investigated. Spatial
108 coherence provides a measure of the potential seasonal predictability, because there is no *a*
109 *proiri* reason for the seasonal anomalies to differ between locations, except due to local scale
110 processes; Moron et al. (2007) argued that these are dominated by unpredictable noise over
111 homogeneous regions. From such analyses of seasonal anomalies, rainfall frequency at local
112 scale has been shown to be generally more spatially coherent in the tropics, and thus
113 potentially more predictable on seasonal time scales (Moron et al. 2007), but it has not
114 heretofore been investigated for the midlatitudes such as is done in this study. For a
115 climatically homogeneous region, even in regions of complex terrain like Coquimbo, high

116 spatial coherence would be an indicator of potential predictability, although the reverse is not
117 necessarily the case.

118 To gain insight into the nature of the daily rainfall variability and its year-to-year
119 modulations in more detail, we model the sequences of station rainfall in terms of different
120 daily rainfall patterns, or ‘rainfall states’, as determined by a hidden Markov model (HMM).
121 The HMM can simulate stochastic daily sequences of rainfall occurrence with a specific
122 rainfall intensity, by estimating the transition probabilities between daily weather patterns or
123 states. The Markov property requires that the probability of occurrence of a particular state on
124 a given day only depends on the previous day’s state. The set of states needed to describe the
125 local daily rainfall characteristics are determined from observed rainfall records; the states are
126 not directly observed and are as such ‘hidden’. In the *homogeneous* HMM, the transition
127 probabilities from one state to the other are not allowed to vary in time. In its *non-*
128 *homogeneous* form (nHMM), the transition probability between states can vary in time,
129 allowing external inputs to influence the rainfall characteristics between one year and another.
130 Seasonal GCM predictions can be used to determine these inputs, creating an effective
131 method to downscale them to most probable daily rainfall sequences at the station scale,
132 training the nHMM on each year for which GCM seasonal hindcasts are available (Charles et
133 al. 1999; Robertson et al. 2004, 2006, 2009). Encouraging results were reported by Bellone et
134 al. (2000), who used a combined climate index, including wind, temperature and relative
135 humidity fields, together with a nHMM to construct a model for daily rainfall amounts. In
136 addition to probabilistic downscaling using the nHMM, we apply a simpler method based on
137 canonical correlation analysis (CCA) to the seasonally-averaged statistics themselves
138 (seasonal amount, daily rainfall frequency, and mean daily intensity), in order to obtain

139 downscaled estimates of their seasonal predictability.

140 The work in this paper aims to lay the foundations for constructing (meteorological)
141 drought early warning systems, through analyses of daily rainfall and by estimating seasonal
142 predictability. The rainfall data and GCMs are described in Sect. 2, with the statistical
143 methods outlined in Sect. 3. The results of the station rainfall analyses and retrospective
144 forecasts of drought indices are presented in Sect. 4, with the concluding remarks in Sect. 5.

145 **2. Data**

146 *a. Observed rainfall data*

147 The analysis of daily rainfall was based on data from 42 stations in the Coquimbo Region,
148 obtained from the Chilean Water Authority (DGA) covering the period 1937–2006. However,
149 it should be noted that data series were only available for a limited number of stations (<13)
150 during the first part (1937–1958) of this period, reaching 38 stations from 1990 onwards
151 (Fig.1).

152 Figures 2 and 3 show the seasonality and spatial distribution of rainfall, together with a
153 decomposition of seasonal rainfall amount into the frequency of occurrence of daily rainfall,
154 and the mean daily intensity of rainfall on wet days (>1 mm). A distinct wet season covering
155 the period May-August (MJJA) accounting for 85% of the annual rainfall amount was
156 identified in the data set (Fig. 2) and was used for further analysis. Rainfall intensities were on
157 average higher during the wet season, with maximum daily rainfall amounts observed
158 between 100 and 207 mm in 22% of the years. In terms of the seasonality within the MJJA
159 season, average rainfall intensities show little within-season systematic modulation, while
160 rainfall frequency shows a clear seasonal modulation with a peak in July.

161 Spatial rainfall characteristics in the Coquimbo region for this period are given in Fig. 3,
162 indicating clear geographical modulation of rainfall. Seasonal rainfall amounts range between
163 43 mm in the north and 270 mm in the pre-Andean cordillera respectively (at 840 m above sea
164 level), tending to increase eastward towards the Andes, due to orographic effects, and from
165 north to south, due to an increased influence of the midlatitude storm track. Rainfall frequency
166 is very low, ranging from 4 to 13 wet days per season on average, and is more geographically
167 modulated than mean rainfall intensity (range of 10 – 24 mm per day). The larger spatial
168 variation of rainfall frequency compared to mean intensity is consistent with the smaller
169 within-season monthly modulation of the latter in Fig. 2, and with the frontal nature of winter
170 rainfall over the region (Aceituno 1988).

171 *b. Seasonal forecast models*

172 Retrospective seasonal MJJA precipitation forecasts initialized on April 1 were obtained from
173 three GCMs: the European Centre Hamburg Model (ECHAM 4.5) (Roeckner et al. 1996), the
174 National Centers for Environmental Prediction (NCEP) Climate Forecast System (CFS) (Saha
175 et al. 2006) and the Community Climate Model (CCM 3.6) (Hurrell et al. 1998). The ECHAM
176 and CCM are atmospheric GCMs that are both driven with the same constructed-analog (CA)
177 *predictions* of global sea surface temperature (SST) in a two-tiered approach (Li and Goddard
178 2005). The two-tier approach has been used as the basis of the International Research Institute
179 (IRI) operational seasonal forecast system since 1997 (Barnston et al. 2010), while studies
180 indicate comparable predictive performance of one- and two-tier approaches (Kumar et al.
181 2008). We thus refer to them in the following as ECHAM-CA and CCM-CA respectively.
182 The CFS is a coupled ocean-atmosphere GCM with initialization of the atmosphere, ocean
183 and land-surface conditions through data assimilation. For all models, the ensemble mean

184 (over 24 members for ECHAM and CCM and 15 members for CFS) gridded precipitation was
185 used at a resolution of T62 (~1.9°) for CFS and T42 (~2.8°) for ECHAM-CA and CCM-CA,
186 over the domain 20°–40°S and 65°–85° W. Seasonal MJJA precipitation hindcasts were
187 available for the 1981–2002 period for CCM-CA and ECHAM-CA, and for the 1981–2005
188 period in the case of CFS.

189 **3. Statistical methods**

190 *a. Spatial coherence analysis*

191 Estimates of spatial coherence of interannual rainfall station anomalies are used as indicators
192 of potential seasonal predictability following Moron et al. (2007). The number of spatial
193 degrees of freedom (DOF) gives an empirical estimate of the spatial coherence in terms of
194 empirical orthogonal functions (EOFs), with higher values denoting lower spatial coherence:

$$195 \quad DOF = \frac{M^2}{\sum_{j=1}^M e_j^2} \quad (\text{Eq. 1})$$

196 where e_j are the eigenvalues of the correlation matrix formed from the station seasonal-mean
197 time series and M is the number of stations.

198 A second measure of the spatial coherence of interannual anomalies is given by the
199 interannual variance of the Standardized precipitation Anomaly Index (var[SAI]), which is
200 constructed from the station average of the standardized rainfall anomalies (Katz and Glantz
201 1986):

202
$$\text{var}[SAI_i] = \text{var}\left[\frac{1}{M} \sum_{j=1}^M \frac{(x_{ij} - \bar{x}_j)}{\sigma_j}\right] \quad (\text{Eq. 2})$$

203 where \bar{x}_j is the long-term time mean over $i=1, \dots, N$ years and σ_j is the interannual standard
 204 deviation for station j . The $\text{var}[SAI]$ is a maximum when all stations are perfectly correlated
 205 ($\text{var}[SAI]=1$) and a minimum when the stations are uncorrelated, resulting in a $\text{var}[SAI]=1/M$.

206 *b. Hidden Markov Model (HMM)*

207 A state-based Markovian model was used to model daily rainfall sequences at the 42 stations,
 208 in order to gain insight into the daily rainfall process, and as a means to downscale daily
 209 rainfall sequences (downscaling in space and time). We use the approach developed by
 210 Hughes and Guttorp (1994) for rainfall occurrence, while additionally modeling rainfall
 211 amounts. The hidden Markov model used here is described fully in Robertson et al. (2004,
 212 2006). In brief, the time sequence of daily rainfall measurements on the network of stations is
 213 assumed to be generated by a first-order Markov chain of a few discrete hidden (i.e.
 214 unobserved) rainfall “states”. For each state, the daily rainfall amount at each station is
 215 modelled here by a delta function at zero amount to model dry days, and an exponential to
 216 describe rainfall amounts on days with nonzero rainfall. To apply the HMM to downscaling
 217 rainfall state transition probabilities were allowed to vary with time, resulting in the
 218 nonhomogeneous HMM (nHMM). In this study, transition probabilities between states are
 219 modelled as functions of predictor variables, in our case GCM predictions of MJJA seasonal-
 220 average precipitation over the region $[5^\circ\text{S}–40^\circ\text{S}, 100^\circ\text{W}–50^\circ\text{W}]$. For data compression, a
 221 conventional principal components (PC) analysis was first applied to the gridded seasonal-
 222 averaged GCM precipitation fields, with each gridded precipitation value standardized by its

223 interannual standard deviation at that gridpoint, selecting here the leading PC as the input
224 variable to the nHMM. The nHMM was trained under leave-three-years-out cross-validation,
225 using the CFS 15-member ensemble mean. To make downscaled simulations, we used each
226 CFS ensemble member in turn, and generated 10 stochastic realizations for each one, yielding
227 an ensemble prediction of 150 daily rainfall sequences for each MJJA season, providing a
228 probabilistic forecast (Robertson et al. 2009).

229 The seasonal statistics of interest (seasonal amount, daily rainfall frequency, and mean daily
230 intensity on wet days) were then computed from these simulated rainfall sequences and
231 compared to their observed counterparts at each station.

232 *c. Downscaling of seasonal forecasts using canonical correlation analysis (CCA)*

233 In addition to the nHMM, downscaling was also carried out by applying canonical correlation
234 analysis (CCA) directly to the seasonal rainfall statistics of interest. The CCA regularizes the
235 high-dimensional regression problem between a spatial field of predictors and predictands by
236 reducing the spatial dimensionality via principal component (PC) analysis and thus minimizes
237 problems of overfitting and multicollinearity (Tippett et al. 2003). Cross validation was used to
238 determine the truncation points of the PC and CCA time series, via the Climate Predictability
239 Tool (CPT) software toolbox (<http://iri.columbia.edu/outreach/software/>). As predictor data
240 sets the retrospective seasonal MJJA precipitation forecasts by the three GCMs discussed in
241 section 2b were used in conjunction with the MJJA seasonal rainfall at the 42 stations of the
242 Coquimbo Region. The three CCA models were trained and tested over the 1981–2000
243 period, with a cross-correlation window of 5 years (i.e. leaving out two years on either side of
244 the verification value). As employed here, the CCA provides a deterministic *mean* forecast

245 value, in contrast to the nHMM probabilistic ensemble.

246 **4. Results**

247 *a. Spatial coherence of rainfall anomalies*

248 As a first step towards assessing the seasonal predictability of rainfall at local scale, we begin
249 with an analysis of spatial coherence, for each of the three rainfall characteristics: seasonal
250 amount, rainfall frequency, and mean daily intensity. Based on Fig. 3, spatial coherence
251 estimates were made separately for the three provinces (from north to south: Elqui, Limari
252 and Choapa) and for three altitude classes (from west to east: 0–500 m, 500–1500 m and
253 >1500m) (Fig. 3). Table 1 shows both the Degrees of Freedom (DOF) and the variance of the
254 Standardized Anomaly Index (SAI) for these sub data sets. The highest DOF and lowest
255 var[SAI] was observed for the Elqui Province in the north, indicating lowest spatial coherence
256 of seasonal anomalies, consistent with its more arid nature and more sporadic rainfall. A
257 similar tendency was found when looking at altitude influences, with lowest spatial coherence
258 at highest altitudes (>1500m), due to orographic influences on rainfall variability.

259 The dependence of spatial coherence characteristics was analysed as a function of time scale
260 using station autocorrelation. Fig. 4 shows the averaged Pearson correlation between each
261 station pair plotted against distance, for rainfall amount, intensity and frequency averaged
262 over several different time windows from daily to seasonal. Taking a value of $1/e$ (0.37) as the
263 decorrelation distance (Dai et al. 1997; Moron et al. 2007; New et al. 2000; Ricciardulli and
264 Sardeshmukh 2002; Smith et al. 2005), anomalies of rainfall amount were found to be
265 significantly correlated at all temporal scales for all stations in the region. A similar
266 observation can be made for rainfall intensity. Rainfall frequency is uncorrelated beyond 150

267 km for the daily and 2-daily time scales, but becomes more highly correlated on longer time
268 scales. This increase in spatial correlation on longer time scales is also found for rainfall
269 amount, but not for intensity. Similar findings were reported by Moron et al. (2007) for
270 tropical rainfall, where it was argued that this increase toward the seasonal scale indicates a
271 common regional seasonal climate forcing on these two rainfall characteristics. Thus, while
272 the occurrence and amount of rainfall at individual stations contain a random element on any
273 particular day, this locally-random element becomes averaged out in time because (a) the
274 atmospheric synoptic storms that impact the region are large-scale and tend to persist over
275 several days (Figs. 6 and 7 below) impacting most stations, and (b) the large-scale impact of
276 ENSO is at the seasonal scale (Sect. 4c below). The spatial autocorrelation function is near-
277 linear and time integration makes the stations more coherent, indicating that the seasonal
278 function is a superposition of daily and seasonal effects, and also that the daily rainfall shows
279 an organized, regional pattern repeated across the season.

280 In contrast to rainfall occurrences and amounts, rainfall intensity does not exhibit any increase
281 in coherence when integrated over time and appears just as coherent at the daily scale. A
282 distance of 200 km could be identified as the decorrelation distance for rainfall intensities
283 between two stations in the Region. This is much larger than found by Moron et al. (2007) in
284 the tropics, but is consistent with the advective character of rainfall in the Coquimbo Region,
285 associated with extratropical cyclones with large spatial scales (Montecinos and Aceituno
286 2003). The examination of spatial coherence statistics in this sub-section indicates that
287 seasonal rainfall amounts and frequencies are likely to be more predictable than mean daily
288 rainfall intensities. However, the differences in spatial coherence between these three seasonal
289 quantities is somewhat less than that found by Moron et al. (2007) in tropical regions.

290 *b. Daily rainfall states*

291 Given the high spatial coherence of rainfall in the Region, we next look in more detail at the
292 evolution of daily rainfall by identifying a small set of typical daily rainfall states (or patterns)
293 and the transitions between them from day to day, using a hidden Markov model. To identify
294 an appropriate number of rainfall states, the log-likelihood of HMMs was computed under
295 cross-validation with up to 10 states, resulting in an increase in log-likelihood for a small
296 number of states, levelling off at higher numbers. This is typical because the rainfall process
297 in nature is more complex than the simple HMM, so that models with more parameters fit the
298 observed rainfall data better, even under cross-validation. For diagnostic purposes, however,
299 we seek a model with a small number of states for interpretability, and a model based on four
300 rainfall states was thus chosen as a compromise. Using the maximum likelihood approach, the
301 HMM parameters were estimated from the entire data set of 7872 days, measured at the 42
302 rainfall stations, applying the iterative expectation-maximization (EM) algorithm (Dempster
303 et al. 1977; Ghahramani 2001). The algorithm was initialized 10 times from random seeds,
304 selecting the run with the highest log-likelihood.

305 The four rainfall states thus obtained are shown in Fig. 5 in terms of their rainfall, showing the
306 probability of rainfall occurrence at each station (panels a–d), and the average rainfall
307 intensity on wet days (panels e–h). States were ordered from overall driest to wettest, showing
308 a dry state 1 with rainfall probabilities near zero at all stations, and three states with increasing
309 probabilities for rainfall and generally larger rainfall amounts on wet days. The spatial pattern
310 of State 2 resembles that of the mean characteristics seen in Fig. 3, with more-frequent rainfall
311 in the south and at higher altitudes. State 3 represents the rainfall events where rainfall is
312 probable at most locations excluding the most northern ones, while rainfall intensities remain

313 relatively small. State 4 can be interpreted as the ‘very wet’ state, with high rainfall
314 probabilities over the whole region, and large rainfall intensities.

315 When looking at the matrix of day-to-day transition probabilities between the four states
316 (Table 2), it can quickly be seen that state 1 is the most persistent state, but it is also the state
317 to which the wetter states 2 and 3 are most likely to evolve. State 4, the very wet state, has an
318 almost equal probability for each of the states to follow it, indicating that states 2 and 3 tend
319 to be intermediate in the transitions from a wet period to a dry period.

320 A visual interpretation of the temporal evolution is given in Fig. 6, showing the most probable
321 daily sequence of the four states that occurred over the 70-winter record (1937-2006) of daily
322 rainfall, obtained using the dynamical programming ‘Viterbi’ algorithm (Forney 1978). Once
323 the parameters of the HMM have been estimated from the rainfall data, the Viterbi algorithm
324 uses the HMM state parameters in conjunction with the rainfall data to assign each day of the
325 historical record to a particular state. This resulted on average in 105 days per season of state
326 1 (85.4%), 10 days of state 2 (7.7%) and 5 days of state 3 (4.2%). The ‘very wet’ state 4
327 occurred on only 3 days (2.7%) on average during each MJJA season of the 70-year period,
328 but on average 56% of total seasonal rainfall was observed on these days. The horizontal
329 traces in Fig. 6 illustrate graphically the high intermittency of rainfall over the region, with
330 individual rainfall events often lasting several days and being made up of days from several of
331 the wetter states. On average, no obvious seasonality is apparent across the season.

332 Figure 7 shows composite sea-level pressure (SLP) fields from the NCEP-NCAR reanalysis
333 data (Kalnay et al. 1996), made by averaging over the days falling into each state, plotted as
334 an anomaly from the long term MJJA average. The state SLP anomaly patterns demonstrate
335 the well-known relationship between rainfall in central Chile and synoptic-wave

336 disturbances (Falvey and Garreaud, 2007). The wet states 2–4 are associated with a similar
337 wave pattern with an anomalous trough over the Chilean coast extending east of the Andes,
338 but with increasing trough intensity as a function of rainfall, while the dry state 1 (note finer
339 contour interval in Fig. 7a) has the opposite footprint of anomalous anticyclonic conditions
340 over central Chile. The tendency seen in the state sequence for multi-day persistent rainfall
341 events made up of several states (Fig. 6) suggests this anomalous low pressure pattern, once
342 established, often remains approximately stationary while growing and decaying in situ.

343 *c. ENSO influence on seasonal rainfall characteristics*

344 The interannual variability over the Coquimbo Region can be interpreted in terms of the
345 HMM's state sequence, with more instances of the wetter states during wet winters. Before
346 proceeding with that analysis, we first summarize the well-known ENSO influence on the
347 seasonal statistics of rainfall (MJJA amount, rainfall frequency, and mean daily intensity)
348 averaged over all 42 stations. The relationship between seasonal rainfall amounts and ENSO
349 is plotted in Fig. 8 in terms of the Niño 3.4 index. All but one (the year 1984) of the very wet
350 winters ($> 100\text{mm}$ above average rainfall) have been associated with the warm ENSO phase,
351 with all of these ENSO events in their developing phase over the MJJA season. The cold
352 ENSO phase has almost always been associated with below normal rainfall, although several
353 years have less than normal rainfall without strong La Niña characteristics. The Pearson
354 correlation between Niño 3.4 and MJJA rainfall amount for the entire period 1937–2006 is
355 0.57, which is statistically significant at the 99% level according to a two-sided Student test.
356 The Spearman correlation coefficient, that is less sensitive than the Pearson correlation to
357 strong outliers, was lower (0.45), but still significant. When only ENSO years are included, as
358 defined by those years where 50% of the MJJA months were marked as ENSO using the

359 Bivariate ENSO or BEST Index (Smith and Sardeshmukh 2000), the Pearson correlation
360 increases to 0.80, which is indicative for the strength of the ENSO signal in extreme wet or
361 dry years. The Spearman correlation coefficient was 0.75, suggesting only a limited influence
362 of outliers.

363 Table 3 shows the correlations between the observed station-averaged MJJA rainfall amount,
364 frequency and intensity and the cross-validated hindcasts from multiple linear regressions
365 with the Niño 3.4 index averaged over different time periods as a predictor. Correlations are
366 strongest when the Niño 3.4 index is contemporaneous or follows the MJJA season, consistent
367 with the so-called ENSO spring predictability barrier around May; once established during
368 boreal summer, ENSO events tend to persist into the following boreal fall. The hindcasts with
369 the FMAM averaged Niño 3.4 index or even for individual months March, April and May,
370 were only weakly correlated with the MJJA total rainfall data, with Pearson correlations of
371 0.26, 0.15, 0.32 and 0.44 respectively, limiting the prediction potential of the Niño 3.4 index.
372 Similar behaviour was found for rainfall frequency, with highest Pearson correlation for the
373 contemporaneous period, whereas rainfall intensity was weakly correlated when using the
374 ENSO index as a predictor for all periods considered (Table 3).

375 *d. ENSO influence on rainfall states*

376 Year-to-year variations in the frequency of the four rainfall states were correlated with the
377 MJJA-average Niño 3.4 index, resulting in Pearson correlation coefficients of -0.44, 0.27,
378 0.18 and 0.52 respectively (all are significant at the 95% level, except for state 3). Thus the
379 ENSO relationship discussed above is mostly expressed in terms of the frequency of
380 occurrences of states 1 and 4. This is remarkable, given the small number of days falling into

381 state 4 and its association with the most-intense storms, and demonstrates the strong
382 relationship between El Niño and intense storms in central Chile.

383 El Niño events tend to weaken the subtropical anticyclone and to displace the frontal storms
384 to more northern locations than normal with a blocking of their usual path further to the south
385 (Garreaud and Battisti 1999; Rutllant and Fuenzalida 1991). This is consistent with our
386 finding of a positive correlation between the occurrence of the three wet states and the ENSO
387 index. When evaluating wet years, Rutllant and Fuenzalida (1991) found that a low-pressure
388 zone becomes established over Central Chile and Northwestern Argentina, separating the
389 Pacific anticyclone from the Atlantic high pressure area, which is consistent with the observed
390 atmospheric circulation patterns observed for states 2 to 4 that exhibit an anomalous synoptic
391 trough between 30°–40°S, and a ridge to the south (Fig.7).

392 *e. Seasonal prediction of daily rainfall aggregates*

393 Given the impact of ENSO on Coquimbo-region rainfall documented in the previous sub-
394 sections, we next explore the seasonal predictability of the observed rainfall based on GCM
395 retrospective forecasts. In this subsection we consider the seasonal aggregate scale, using the
396 canonical correlation analysis (CCA) described in Sect 3c, to regress the GCM seasonal-
397 average rainfall predictions onto the observed station seasonal rainfall statistics presented in
398 Sect 4c. Scatterplots of the cross-validated seasonal rainfall deterministic forecasts are shown
399 in Fig. 9 over the hindcast period (1981–2000) for each of the three GCMs, where each circle
400 represents the forecast mean of the seasonal rainfall amount for each station-year. A clear
401 deviation from the 1:1 line is observed for the ECHAM-CA model, indicating clear
402 underestimation of the higher rainfall amounts observed during wet years. The CCM-CA
403 model shows an overestimation at the lower rainfall amounts, while failing to predict the

404 more extreme rainfall values. The CFS model performs best, showing the least scatter as well
405 as quite successful predictions in the higher range of rainfall amounts. This is confirmed by
406 Table 4, which gives the station-averaged root mean square error (RMSE), mean error or bias
407 (ME), and Pearson anomaly correlation coefficients for each of the model (cross-validated)
408 retrospective forecasts of station precipitation. The CCM-CA gave the lowest correlation and
409 the highest RMSE, but was the least biased, with a low ME. Correlation was higher for the
410 ECHAM-CA, but ME and RMSE indicated an important bias in comparison to the other
411 models. The CFS model showed the highest correlation coefficient and a low RMSE, but with
412 a negative ME, underestimating the observed rainfall amounts at the highest observed rainfall
413 amounts (e.g. 36% at 500mm). Nevertheless, the CFS model was selected for further
414 processing, due to its superior correlation statistics. Since seasonal MJJA hindcasts for the
415 period 1981-2005 were available for the CFS model, this period was used for further analysis.

416 The Pearson correlation skill map from CFS for all stations (Fig. 10) shows a good correlation
417 between observed and hindcast precipitation for almost all stations, with individual
418 correlations between 0.57 and 0.80. This could be expected, due to the high skill (R was 0.76)
419 of the CFS to predict Niño 3.4 SST, when initialized on April 1, and a high correlation (R of
420 0.82) between the leading PC of the gridded CFS rainfall and Niño 3.4 SST, that explains
421 large part of the variability in rainfall amounts observed (see Fig. 8). A similar picture
422 emerges for rainfall frequency, with slightly lower Pearson skill (0.20–0.63), while the
423 correlation coefficients for rainfall intensity are generally much lower (from -0.15 until 0.64).

424 *f. Seasonal prediction of stochastic daily rainfall sequences*

425 Having addressed the seasonal predictability of daily rainfall aggregates (seasonal amount,
426 rainfall frequency, and mean daily intensity) in the previous subsection, we next use the

427 nHMM as described in Sect. 3b to derive seasonal forecasts of daily rainfall sequences at each
428 of the stations. The nHMM used here builds on the HMM results presented in Sect 4b and 4d,
429 but with the inclusion of CFS forecasts of MJJA seasonal-average precipitation, as described
430 in Sect. 3b. This is the same CFS predictor field used via CCA in the previous subsection.

431 The resulting daily rainfall simulations were then used to construct the seasonal rainfall
432 amount, rainfall frequency, and mean daily intensity, and the ensemble averages then
433 correlated with observed values (Fig.11). As in the case of the CCA-based forecasts in Fig.
434 10, correlations were generally higher for seasonal rainfall amount and rainfall frequency,
435 compared with rainfall intensity, with station values ranging from 0.17–0.92, 0.19–0.92 and
436 -0.38–0.84 for the three quantities respectively. Inter-station differences in skill are larger than
437 in the CCA approach, but fewer stations with negative correlations were obtained using the
438 nHMM (note that only positive values are plotted in Figs. 10 and 11).

439 Time series of the station-averaged MJJA seasonal rainfall statistics are plotted in Fig. 12,
440 which compares the median and interquartile range of the 150-member ensemble of nHMM
441 simulations, together with the observed values and CCA-based hindcasts. The hindcasts of
442 seasonal rainfall amount obtained using both methods (CCA and nHMM) follow the observed
443 highs and lows reasonably well (Fig. 12a), with a Pearson correlation skill for the CCA of
444 0.77 and for the nHMM mean of 0.62. A small overestimation for the nHMM low rainfall
445 amount years is observed, as well as an underestimation when dealing with very wet years,
446 and can be attributed to the deviations observed between the CFS model predicted and
447 observed precipitation (Table 4 and Fig. 9).

448 The rainfall frequency hindcasts from the nHMM were also skillful ($\rho=0.55$), representing the
449 observed interannual variability better than for rainfall intensity ($\rho=0.30$). The

450 underestimated mean rainfall intensities in 1984 resulted in an important underestimation of
451 the seasonal rainfall amount for nHMM, while the CCA hindcast of seasonal amount was less
452 affected. The year 1983, on the other hand, had more rainfall days than picked up by the
453 nHMM and CCA, but with low intensities, still resulting in acceptable predictions of the
454 seasonal rainfall amount with both methods.

455 *g. Towards a Drought Early Warning System*

456 Although no effort was made to design or setup a drought early warning system for the
457 Coquimbo Region, this paper tries to identify the prediction potential of meteorological
458 drought indices that would be essential to such an effort. In order to tailor our rainfall
459 hindcasts more specifically to drought indicators, we firstly express our hindcasts in terms of
460 the Standardized Precipitation Index or SPI (Edwards and McKee 1997; McKee et al. 1993), a
461 commonly used meteorological drought classification method. The SPI is derived by
462 transforming the probability distribution of (here seasonal amount) rainfall into a unit normal
463 distribution so that the mean SPI is zero and each value is categorized in one of its 5 quantiles
464 and as such given a 'drought class'. The SPI hindcasts derived from the CFS using the CCA
465 and nHMM methods are shown in Fig. 13, together with those derived from the
466 (simultaneous) regression with MJJA Niño 3.4 SST. Each of the methods was able to
467 represent observed SPI variability rather well, although different SPI classes were often
468 predicted. This is reflected in Table 5, where fits between observed and simulated SPIs are
469 expressed in terms of ρ , ME, RMSE and hit rate. None of the prediction models showed a
470 significant bias, due to the standardization of the SPI, and the Pearson correlation values are
471 similar to those of the seasonal rainfall amount. The hit rate measures the success of the
472 method to predict the SPI class, and substantially exceeds the 20% rate expected by chance.

473 When interested in general trends, the hit rate might be too strict to evaluate the prediction
474 skill. For example, for the year 1987 CCA and nHMM predicted SPI values of 2.1 and 2.6
475 respectively ('extremely wet'), whereas the observed SPI value was 1.8 ('very wet'), reducing
476 the hit score although the prediction was rather accurate. When accepting model predictions
477 that are one SPI class lower, equal to or one class higher than the observed SPI class, the hit
478 rate increased to values of 92%, 92% and 88% for the ENSO Index, the CCA method and the
479 nHMM respectively, indicating that the general trend (dry, wet or normal) is maintained by
480 the three methods.

481 In a second approach, the nHMM hindcasts of daily rainfall series were converted into four
482 daily drought indices and compared with observed values. As a first index, the frequency of
483 days with substantial rainfall (>15 mm per day) was used to classify years with drought risk,
484 hereafter named *FREQ15*. A second set of drought indices based on daily rainfall were
485 derived from the work of Byun and Wilhite (1999), using the Effective Precipitation (EP)
486 concept to represent the soil water storage at each day during the wet season. After evaluating
487 all proposed indices, three were retained, based on their higher sensitivity to observed drought
488 in the region. The first is the Accumulated Precipitation Deficit (APD), which gives a simple
489 accumulation of daily precipitation deficit from May to August, and was found to be a good
490 measure for drought intensity. The second is the Accumulation of Consecutive days of
491 Negative Standardized EP, further named *ANES*, and is a measure for accumulated stress
492 during droughts (Byun and Wilhite 1999). The last is the Effective Drought Index or EDI,
493 which gives a drought classification similar to the SPI, with positive values indicating surplus
494 rainfall and negative values dry or drought conditions.

495 In general, the drought indices obtained from the nHMM cross-validated hindcasts closely

496 followed the observed indices for the period 1981–2005. The fit proved best for the FREQ15
497 and the APD index ($\rho=0.63$ and 0.60 resp.), compared with the EDI and ANES ($\rho= 0.51$ and
498 0.50). The observed versus predicted values of FREQ15 and APD are plotted in Fig. 14,
499 which shows the median and the inter-quartile range (IQR) of the nHMM ensemble, showing
500 information about the predicted distributions. The spread of the forecasts can be evaluated in
501 terms of the IQR, which should bracket the observations in 50% of the years to be well
502 calibrated, with lower values indicating too little spread and with values above 50% for those
503 forecast distributions with too much spread. For both drought indices, the inter-quartile range
504 brackets the observed values in 44% and 52% of the years respectively, indicating that the
505 forecasts are rather well calibrated.

506 To quantify the skill of these probabilistic forecasts, the Ranked Probability Skill Score
507 (RPSS) is used, which is a squared error metric that allows measuring the distance between
508 the cumulative distribution function of the forecast and the verifying observation, and is
509 expressed with respect to a baseline given by the climatic forecast distribution. A perfect
510 forecast would be represented by a RPSS of 100%, while negative values indicate that the
511 forecast is less skillful than the climatological equal-odds forecast. For the four indices under
512 consideration, the station average median RPSS values were 25.2 for FREQ15, 12.3 for APD,
513 12.7 for EDI and -2.6 for ANES, indicating better than climatology forecasts in the former
514 three cases. Additionally, the percentage of positive RPSS values were found to be 84%, 80%,
515 64% and 48% respectively, confirming the better predictability of FREQ15 and APD, and
516 lower predictability for EDI and ANES. In Fig 15 the median RPSS is presented for each
517 station for FREQ15 and APD, showing similar results as the station average, with few areal
518 differences, but suggests a superior predictability of FREQ15 compared to APD.

519 Both indices can also be related to declared drought years in the Coquimbo region (Novoa-
520 Quezada 2001), as defined by seasonal rainfall amounts not exceeding a minimum threshold
521 to recharge the topsoil during the wet season (e.g. 207 mm in the Southwestern part of the
522 Coquimbo Region), when evaluated with a regional water balance method. Since both
523 FREQ15 and APD showed a correlation of 0.87 with declared drought years, their relatively
524 good predictability is especially encouraging for climate risk management purposes.

525 **5. Summary and conclusions for dry-land management**

526 Rainfall variability is known to be a major economic and social disruptor in the central-
527 northern area of Chile, with large financial consequences for society when both extreme
528 drought and extreme wet conditions occur. A low preparedness could be partly responsible for
529 the large impact of these events. Therefore, the Chilean government has supported the
530 development of a climate risk management system for the semi-arid regions of Chile, to
531 reduce the vulnerability and increase resistance to extreme climatic events. An early warning
532 system for droughts and floods would be an essential component of such an approach, which
533 requires estimation and prediction of the rainfall characteristics relevant to drought as a first
534 step.

535 Winter rainfall characteristics in the Coquimbo Region of Chile were first investigated
536 using daily rainfall records at 42 stations, with special attention to spatial and temporal
537 characteristics and the relationship with ENSO. Seasonal rainfall amounts, daily rainfall
538 frequencies, and mean daily rainfall intensities all generally increase southward and eastward
539 toward the Andes (Fig. 3). An analysis of the spatial correlations between stations (Fig. 4)
540 indicated large inter-station correlations at the daily timescale, particularly for rainfall
541 amount. The spatial coherence of rainfall amount and frequency was found to increase

542 substantially with temporal averaging, suggesting the role of ENSO forcing at the seasonal
543 scale. Seasonal anomalies of mean daily rainfall intensity were found to be less spatially
544 coherent (Table 1), though their coherence was larger than found by Moron et al. (2007) for
545 tropical rainfall, due to the frontal character of rainfall in the region.

546 The spatio-temporal evolution of daily rainfall patterns across the region was further
547 elucidated in terms of four rainfall “states” identified using a Hidden Markov Model (HMM);
548 these states consisted of dry and increasingly wet conditions (Fig. 5), the latter associated with
549 near-stationary trough in sea-level pressure, centered to the south and east of the region (Fig.
550 7). The daily sequences of these states showed sporadic rainfall events with little seasonality
551 within the winter season (Fig. 6), while the likelihood of an intense storm across the region
552 (state 4) was found to be strongly correlated with ENSO, thus providing an interpretation in
553 terms of daily weather for the well-established seasonal rainfall relationship with ENSO (Fig.
554 8).

555 Seasonal predictability of rainfall characteristics was explored firstly using a simple univariate
556 index of ENSO; this proved only to be well correlated for simultaneous (May or MJJA)
557 values of the index, and thus not useful for prediction since lead times are insufficient for
558 drought prediction. Rainfall intensities were found not to be well correlated. Predictability
559 was further explored using a GCM to forecast MJJA rainfall amounts. In our approach, the
560 GCMs were initialized with April 1 climate and/or oceanic conditions of each year 1981–
561 2005, presenting as such a real prediction with lead times up to four months. Of three the
562 GCMs considered (ECHAM-CA, CCM-CA and CFS), the highest skill and lowest bias was
563 obtained for the CFS model (Fig. 9; Table 4). The CFS was then downscaled to represent
564 local variability in station data, using two different techniques. First, a Canonical Correlation

565 Analysis (CCA) approach was developed to map GCM seasonal forecasts of seasonal
566 precipitation to seasonal rainfall characteristics (seasonal rainfall amount, daily rainfall
567 frequency, mean daily rainfall intensity on wet days) at each rainfall station. Secondly, a non-
568 homogeneous Hidden Markov Model (nHMM) was used to derive ensembles of stochastic
569 daily rainfall sequences at each station as a function of GCM seasonally averaged rainfall; the
570 seasonal rainfall characteristics were then calculated from these simulated daily sequences.
571 For both downscaling methods the skill for seasonal rainfall amount, frequency, and mean
572 daily intensity examined at the station scale (Figs. 10–12) produced similar results. The
573 highest correlations with observations were found for seasonal rainfall amount and rainfall
574 frequency for most measuring stations in the region, but low or negative correlations for
575 rainfall intensity. These differences in skill are consistent with differences in the spatial
576 coherence of station-scale seasonal rainfall anomalies (Table 1), with mean daily rainfall
577 intensity being less spatially coherent and thus less predictable than seasonal amount and
578 daily frequency (Moron et al. 2007).

579 Since the objective of the work is oriented towards the development of a drought early
580 warning system for dry-land management, the (retrospective) forecasts of rainfall were then
581 tailored for drought prediction. Following the recommendations of the World Meteorological
582 Organization (Declaration on Drought Indices, December 11, 2009), the Standardized
583 Precipitation Index (SPI), which was calculated from seasonal rainfall amounts, was used as a
584 proxy for meteorological drought. The SPI was forecast with both the CFS-CCA and CFS-
585 nHMM approach and compared with observed values, showing that the SPI was quite well
586 forecast by both methods (Fig. 13).

587 For some end-user applications, the seasonal-average SPI may be too coarse, and drought

588 indices based on daily weather statistics may be more appropriate. Motivated by the potential
589 needs of dry-land management, the nHMM was used to forecast four additional drought
590 indices based on daily rainfall statistics, of which the frequency of occurrence of days
591 exceeding 15 mm/day (FREQ15) and the accumulated daily precipitation deficit (APD) gave
592 highest correlations with observations and positive prediction skill for all stations. While the
593 CCA could also be applied to these statistics, calculating the appropriate predictand from
594 daily observed data, the daily rainfall sequences simulated by the nHMM have the potential to
595 be used in pasture and crop models which require daily weather sequences (e.g. Robertson et
596 al. 2007).

597 Downscaled seasonal predictions of seasonal and daily rainfall characteristics and related
598 meteorological drought indices have been shown feasible for the Coquimbo Region. This
599 could be regarded as an important step in the development of a tailored climate risk
600 management system that should contribute to reduce climate uncertainty in a region that is
601 affected by high rainfall variability. The approach presented in this paper could eventually be
602 extended to forecast agricultural and/or hydrological drought conditions, for which high
603 spatial and temporal resolution of downscaled predictions is required, such as provided by the
604 nHMM approach. The methodology for predicting the nature of within-season daily rainfall
605 variability presented in this paper is also likely to be successful in other regions where daily
606 rainfall variability can be linked to predictable large-scale climatic patterns.

607 **ACKNOWLEDGEMENTS**

608 The authors would like to thank the Chilean Water Authority (Dirección General de
609 Aguas) for providing the daily rainfall series of the Coquimbo Region. This research was
610 funded by the Flemish Government, Department of Sciences and Innovation/Foreign Policy,
611 the UNESCO Regional Office for Science and Technology in Latin America and the
612 Caribbean, and by the US National Oceanic and Atmospheric Administration through a block
613 grant to the IRI.

614 **References**

- 615 Aceituno, P., 1988: On the functioning of the Southern Oscillation in the South American
616 sector. Part I: Surface climate. *Mon. Wea.Rev.*, 505-524.
- 617 Aceituno, P., M. Prieto, M. Solari, A. Martínez, G. Poveda, and M. Falvey, 2009: The 1877–
618 1878 El Niño episode: associated impacts in South America. *Climatic Change*, **92**, 389-416.
- 619 Barnston, A. G., S. Li, S. J. Mason, D. G. DeWitt, L. Goddard, and X. Gong, 2010:
620 Verification of the first 11 years of IRI's seasonal climate forecasts. *J. Appl. Meteor. Climat.*,
621 **49**, in press.
- 622 Bellone, E., J. P. Hughes, and P. Guttorp, 2000: A hidden Markov model for downscaling
623 synoptic atmospheric patterns to precipitation amounts. *Climate Res.*, **15**, 1-12.
- 624 Byun, H.-R., and D. A. Wilhite, 1999: Objective Quantification of Drought Severity and
625 Duration. *J. Climate*, **12**, 2747-2756.
- 626 Charles, S. P., B. C. Bates, P. H. Whetton, and J. P. Hughes, 1999: Validation of downscaling
627 models for changed climate conditions: case study of southwestern Australia. *Clim. Res.*, **12**,
628 1-14.
- 629 Dai, A., I. Y. Fung, and A. D. del Genio, 1997: Surface global observed land precipitations
630 variations: 1900-1988. *J. Climate*, **10**, 2943-2962.
- 631 Dempster, A. P., N. M. Laird, and D. R. Rubin, 1977: Maximum likelihood from incomplete
632 data via the EM algorithm. *J. R. Stat. Soc. B.*, **39**, 1-38.

633 Edwards, D. C., and T. B. McKee, 1997: Characteristics of 20th century drought in the United
634 States at multiple time scales, Climatology Rep. 97-2, Department of Atmospheric Science,
635 Colorado State University, Fort Collins, Colorado, 155 pp.

636 Falvey, M. and R. Garreaud, 2007: Wintertime Precipitation Episodes in Central Chile:
637 Associated Meteorological Conditions and Orographic Influences. *J. Hydrometeorol.*, **8**, 171-
638 193.

639 Forney, G. D., 1978: The Viterbi algorithm. *Proc. IEEE*, **61**, 268-278.

640 FOSIS, 2008: Superando la sequía. Plan Especial para familias vulnerables de zonas rurales
641 afectadas por la sequía, 11 pp.

642 Garreaud, R. D., and D. S. Battisti, 1999: Interannual (ENSO) and interdecadal (ENSO-like)
643 variability in the Southern Hemisphere tropospheric circulation. *J. Climate*, **12**, 2113-2122.

644 Garreaud, R. D., M. Vuille, R. Compagnucci, and J. Marengo, 2009: Present-day South
645 American climate. *Palaeogeogr. Palaeocl.*, **281**, 180-195.

646 Ghahramani, Z., 2001: An introduction to hidden Markov models and Bayesian networks. *Int.*
647 *J. Pattern Recogn. Artif. Intell.*, **15**, 9-42.

648 Goddard, L., A. G. Barnston, and S. J. Mason, 2003: Evaluation of the IRI's "Net Assessment"
649 seasonal climate forecasts: 1997-2001. *Bull. Amer. Meteor. Soc.*, **84**, 1761-1781.

650 Hughes, J. P., and P. Guttorp, 1994: A class of stochastic models for relating synoptic
651 atmospheric patterns to regional hydrologic phenomena. *Water Resour. Res.*, **30**, 1535-1546.

652 Hurrell, J. W., J. J. Hack, B. A. Boville, D. L. Williamson, and J. T. Kiehl, 1998: The

653 dynamical simulation of the NCAR Community Climate Model version 3 (CCM3). *J.*
654 *Climate*, **11**, 1207-1236.

655 Kalnay, E., and Coauthors, 1996: The NCEP/NCAR 40-Year Reanalysis Project. *Bull. Amer.*
656 *Meteor. Soc.*, **77**, 437-471.

657 Katz, R. W., and M. H. Glantz, 1986: Anatomy of a rainfall index. *Mon. Wea. Rev.*, **114**, 764-
658 771.

659 Kumar, A., Q. Zhang, J.-K. E. Schemm, M. Heureux, and K.-H. Seo, 2008: An Assessment of
660 Errors in the Simulation of Atmospheric Interannual Variability in Uncoupled AGCM
661 Simulations. *J. Climate*, **21**, 2204-2217.

662 Li, S., and L. Goddard, 2005: Retrospective Forecasts with the ECHAM4.5 AGCM. IRI
663 Technical Report 05-02. IRI, Palisades, NY.

664 McKee, T. B., N. J. Doesken, and J. Kliest, 1993: The relationship of drought frequency and
665 duration to time scales. *Proceedings of the 8th Conference on Applied Climatology*, A. M.
666 Society, Ed., 179-184.

667 Montecinos, A., and P. Aceituno, 2003: Seasonality of the ENSO-Related Rainfall Variability
668 in Central Chile and Associated Circulation Anomalies. *J. Climate*, **16**, 281-296.

669 Moron, V., A. W. Robertson, M. N. Ward, and P. Camberlin, 2007: Spatial coherence of
670 tropical rainfall at the regional scale. *J. Climate*, **20**, 5244-5263.

671 New, M., M. Hulme, and P. Jones, 2000: Representing twentiethcentury space-time climate
672 variability. Part II: Development of 1901-96 monthly grids of terrestrial surface climate. *J.*
673 *Climate*, **13**, 2217-2238.

674 Novoa-Quezada, P., 2001: Determinación de condición de sequía por análisis de la variación
675 de humedad del horizonte radicular usando el modelo de simulación Hidrosuelo. CONAF,
676 Viña del Mar, Chile, 44 pp.

677 Pittock, A. B., 1980: Patterns of climatic variation in Argentina and Chile. Part I:
678 Precipitation, 1931-1960. *Mon. Wea. Rev.*, **108**, 1347-1361.

679 Quinn, W., and V. Neal, 1983: Long-term variations in the Southern Oscillation, El Niño and
680 the Chilean subtropical rainfall. *Fish.Bull.*, **81**, 363-374.

681 Ricciardulli, L., and P. D. Sardeshmukh, 2002: Local time- and space scales of organized
682 tropical deep convection. *J. Climate*, **15**, 2775-2790.

683 Robertson, A. W., S. Kirshner, and P. J. Smyth, 2004: Downscaling of daily rainfall
684 occurrence over northeast Brazil using a hidden Markov model. *J. Climate*, **17**, 4407-4424.

685 Robertson, A. W., A. V. M. Ines, and J. W. Hansen, 2007: Downscaling of Seasonal
686 Precipitation for Crop Simulation. *J. App Meteorol Clim*, **46**, 677-693.

687 Robertson, A. W., V. Moron, and Y. Swarinoto, 2009: Seasonal predictability of daily rainfall
688 statistics over Indramayu district, Indonesia. *Int. J. Climatol.*, **29**, 1449-1462.

689 Robertson, A. W., S. Kirshner, P. C. Smyth, S. P. Charles, and B. C. Bates, 2006:
690 Subseasonal-to-interdecadal variability of the Australian monsoon over North Queensland. *Q.*
691 *J. R. Meteorol. Soc.*, **132**, 519-542.

692 Roeckner, E., and Coauthors, 1996: The atmospheric general circulation model ECHAM-4:
693 Model description and simulation of present-day climate, 90 pp.

694 Rubin, M. J., 1955: An analysis of pressure anomalies in the Southern Hemisphere. *Notos*, **4**,
695 11-16.

696 Rutllant, J., and H. Fuenzalida, 1991: Synoptic aspects of the central Chile rainfall variability
697 associated with the Southern Oscillation. *Int. J. Climatol.*, **11**, 63-76.

698 Rutllant, J., 2004: Large-scale atmospheric circulation features associated with the 1997-1999
699 ENSO cycle and their consequences in the central-Chile precipitation regime. *El Niño-La*
700 *Niña 1997-2000. Its effects in Chile*, S. Avaria, J. Carrasco, J. Rutllant, and E. Yáñez., Eds.,
701 CONA, 61-76.

702 Saha, S., and Coauthors, 2006: The NCEP Climate Forecast System. *Journal of Climate*, **19**,
703 3483-3517.

704 Smith, C. A., and P. Sardeshmukh, 2000: The Effect of ENSO on the Intraseasonal Variance
705 of Surface Temperature in Winter. *Int. J. Climatol.*, **20**, 1543-1557.

706 Smith, D. F., A. J. Gasiewski, D. L. Jackson, and G. A. Wick, 2005: Spatial scales of tropical
707 precipitation inferred from TRMM Microwave imager data. *IEEE Trans. Geophys. Remote*
708 *Sens.*, **43**, 1542-1551.

709 Tippett, M. K., M. Barlow, and B. Lyon, 2003: Statistical correction of central Southwest
710 Asia winter precipitation simulations. *Int. J. Climatol.*, **23**, 1421-1433.

711 Table 1 Degrees of freedom (DOF) and variance of the Standardized Anomaly Index
 712 (Var[SAI]) for seasonal rainfall amount (RAm), rainfall intensity (RI) and rainfall frequency
 713 (RF) for each province and altitude class in the Coquimbo Region

	N [†]	DOF			Var[SAI]		
		RAm	RI	RF	RAm	RI	RF
Province							
Elqui	12	9.14	17.26	11.85	0.30	0.20	0.26
Limari	24	4.07	7.13	5.10	0.44	0.31	0.40
Choapa	8	5.14	7.17	6.18	0.41	0.34	0.37
Altitude class							
0–500 m	13	5.05	8.91	6.33	0.40	0.28	0.36
500–1500 m	25	5.10	7.86	6.38	0.40	0.31	0.35
>1500 m	6	13.57	21.01	15.09	0.23	0.16 [‡]	0.21

714 [†] Number of stations used

715 [‡] Minimum Var[SAI] value (mean correlation equals zero)

716 Table 2 Transition matrix for the 4-state HMM. “From” states occupy the rows, “To” states

717 the columns. Thus, the probability of a transition from state 2 to state 1 is 0.56.

From State	To State			
	1	2	3	4
1	0.92	0.05	0.02	0.01
2	0.56	0.22	0.12	0.09
3	0.47	0.25	0.19	0.09
4	0.23	0.28	0.22	0.27

718

719 Table 3 Pearson correlation coefficients between the observed station-averaged seasonal
 720 rainfall amounts (RAm), rainfall frequencies (RF) and rainfall intensities (RI), and the cross-
 721 validated hindcasts using the average Niño 3.4 index (1937-2006) for different months and
 722 multi month periods as a predictor.

	Average Niño 3.4 Index					
	FMAM	March	April	May	MJJA	NDJF
RAm	0.26	0.15	0.32	0.44	0.57	0.39
RF	0.33	0.23	0.36	0.49	0.59	0.39
RI	-0.29	-0.42	-0.09	0.03	0.20	0.19

723

724 Table 4 Pearson correlation coefficient (ρ), mean error (ME) and root mean squared error
725 (RMSE) for cross-validated CCA hindcasts of seasonal rainfall amount with the ECHAM-CA,
726 CFS and CCM-CA models for the period 1981-2000.

	ρ (-)	ME (mm)	RMSE (mm)
ECHAM-CA	0.41*	-9.74	133.54
CFS	0.69*	-8.45	99.92
CCM-CA	0.17*	-2.24	137.73

727 * Correlation is significant at $\alpha=0.05$

728 Table 5 Pearson correlation coefficient (ρ), mean error (ME), root mean squared error
 729 (RMSE) and hit rate for hindcasts expressed in terms of the station-averaged Standardized
 730 Precipitation Index. The hindcasts were made based on the MJJA Niño 3.4 Index, and the
 731 CFS downscaled with CCA or the nHMM.

	Niño 3.4 Index	CCA	nHMM
ρ	0.56	0.65	0.58
ME	-0.01	-0.01	0.00
RMSE	0.17	0.17	0.18
Hit Rate [†] (%)	68.0	56.0	60.0

732 [†] Defined as the percentage of correct SPI class prediction

733 **List of Figures**

734 FIG. 1. Number of stations with May–August station rainfall data used in the analysis.

735 FIG. 2. Box plots of rainfall seasonality in the Coquimbo Region; a) number of wet days
736 ($>1\text{mm}$) and b) rainfall intensity (mm day^{-1}) on wet days.

737 FIG. 3. Average rainfall characteristics during the wet season (May–August): a) seasonal
738 rainfall amount; b) rainfall frequency; and c) mean daily rainfall intensity for the period 1937–
739 2006. A locator map indicates the position of the Coquimbo Region within Chile and South
740 America.

741 FIG. 4. Spatial correlation function of a) daily rainfall amount, b) rainfall intensity and c)
742 rainfall frequency for the daily, 2-day, 5-day, 10-day, 30 day and seasonal averaging periods.
743 The $1/e$ decorrelation value is shown as a dashed line.

744 FIG. 5. Four-state HMM rainfall parameters. (a)–(d): probabilities of rainfall occurrence and
745 (e)–(h): mean rainfall intensities (i.e. wet-day amounts).

746 FIG. 6. The most probable HMM state sequence obtained using the Viterbi algorithm. Rainfall
747 states are indicated from driest (state 1) to wettest (state 4) on the grey color bar.

748 FIG. 7. Composites of sea-level pressure anomalies (mb) for each rainfall state. A finer
749 contour interval was used in panel (a) for clarity.

750 FIG. 8. Station-average MJJA rainfall amount, colored according to the sign and magnitude of
751 the Niño 3.4 SST index for the period 1937–2005.

752 FIG. 9. Cross-validated hindcasts versus observed precipitation amounts using CCA for the
753 three GCMs for the period 1981–2000, where each circle represents the value for each station,
754 for each year. Thus there are 42x20 circles in each panel.

755 FIG. 10. Pearson correlation between CFS hindcasts downscaled using CCA and observed
756 rainfall for (a) seasonal rainfall amount, (b) rainfall frequency, and (c) mean rainfall intensity,
757 for the period 1981–2005.

758 FIG. 11. As Fig. 10, but for CFS downscaled rainfall obtained using the nHMM and taking the
759 ensemble mean over the 150 nHMM simulations.

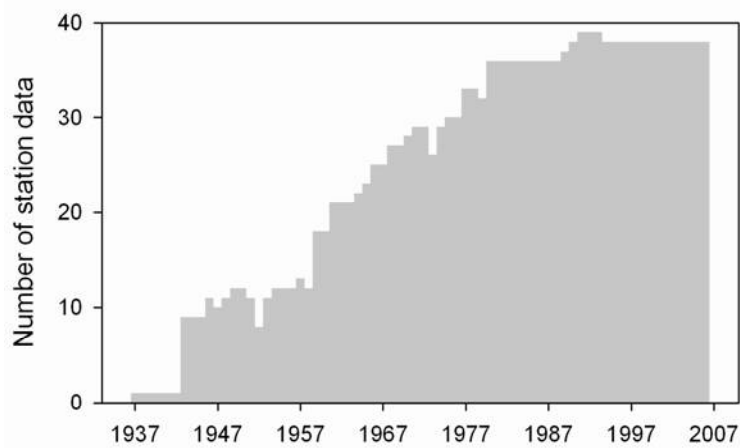
760 FIG. 12. Comparison of station-averaged downscaling results obtained from CFS using CCA
761 and the nHMM. (a) seasonal rainfall amount, (b) rainfall frequency (c), mean rainfall
762 intensity. The error bars indicate the 25th and 75th percentiles of the simulated nHMM
763 ensemble values. The Niño 3.4 index is also indicated.

764 FIG. 13. Station-averaged hindcasts of the Standardized Precipitation Index (SPI) obtained
765 from the MJJA Niño 3.4 Index, and the CFS downscaled with CCA and with the nHMM. The
766 SPI values constructed from observed station rainfall are also plotted.

767 FIG. 14. Hindcasts of two station-averaged meteorological drought indices (circles) consisting
768 of (a) the number of days with rainfall exceeding 15mm (FREQ15), and (b) the Accumulated
769 (daily) Precipitation Deficit (APD), based on the nHMM simulations, compared with values
770 constructed from observed daily rainfall (dashed). The error bars indicate the 25th and 75th
771 percentiles of the nHMM ensemble. The Niño 3.4 index is also indicated.

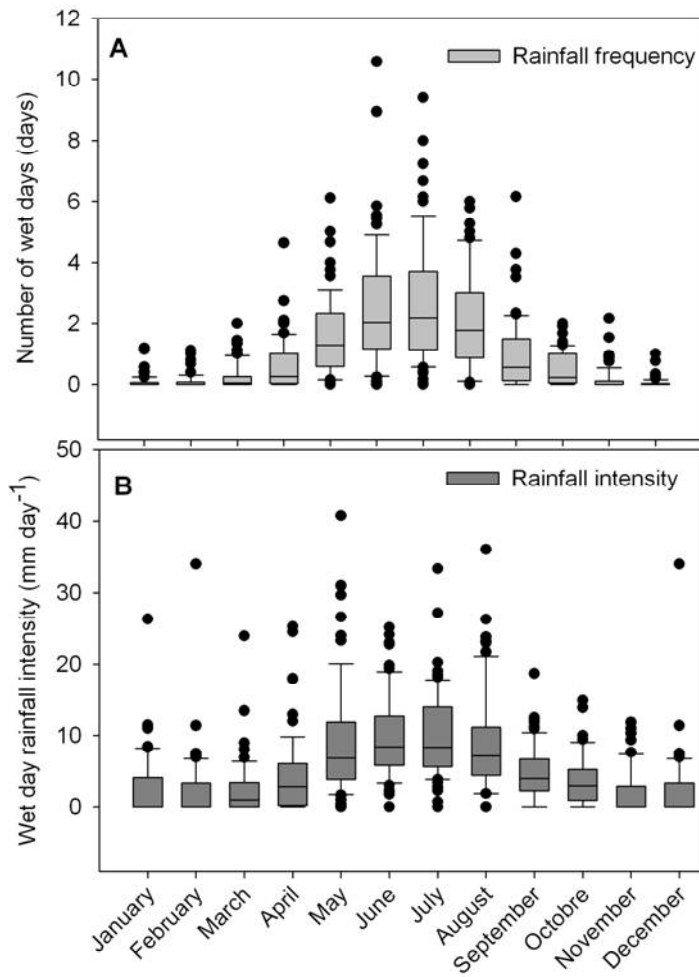
772

773 FIG. 15. Median Rank Probability Skill Score (RPSS) for hindcasts of a) the number of days
774 with rainfall exceeding 15mm (FREQ15) and b) the Accumulated Precipitation Deficit
775 (APD), constructed from the nHMM simulations.



776

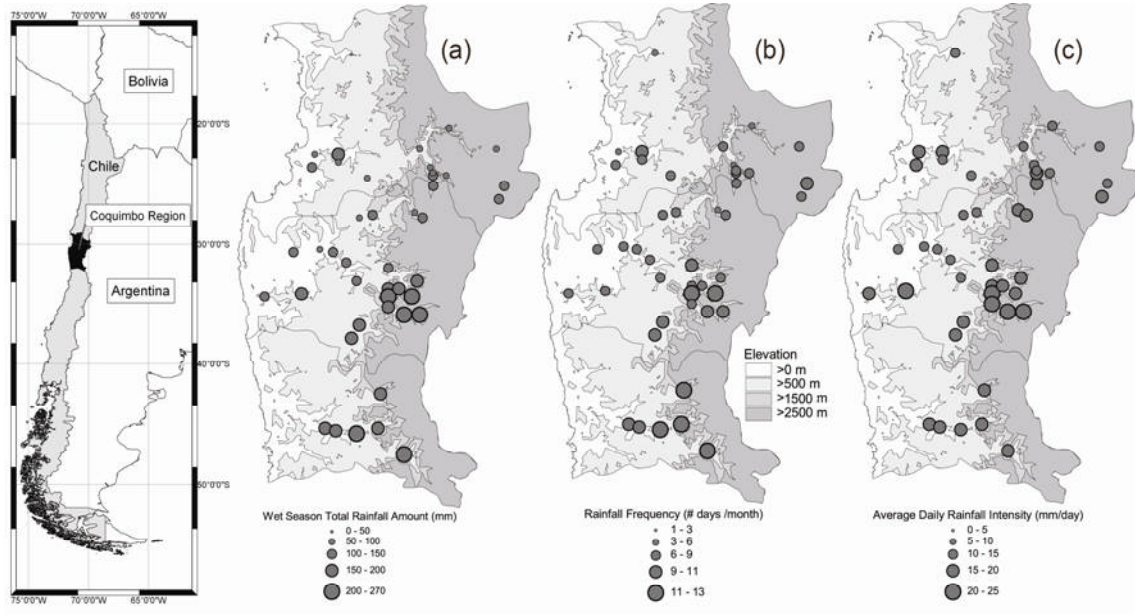
777 FIG. 1. Number of stations with May–August station rainfall data used in the analysis.



778

779 FIG. 2. Box plots of rainfall seasonality in the Coquimbo Region; a) number of wet days
 780 (>1mm) and b) rainfall intensity (mm day⁻¹) on wet days.

781



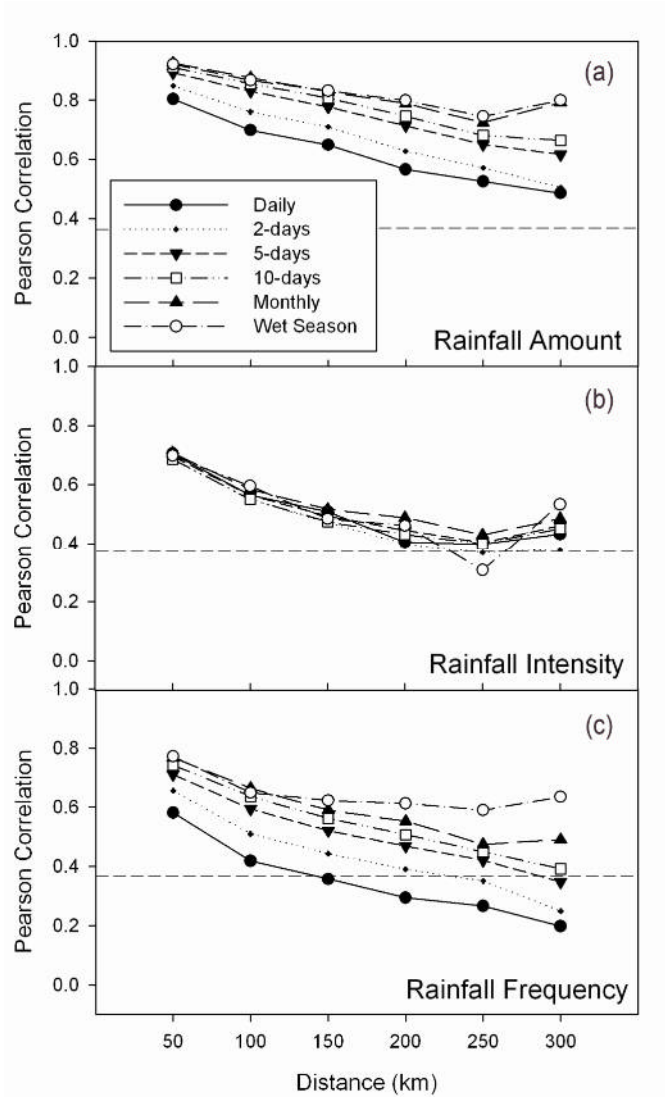
782

783 FIG. 3. Average rainfall characteristics during the wet season (May–August): a) seasonal
 784 rainfall amount; b) rainfall frequency; and c) mean daily rainfall intensity for the period 1937–
 785 2006. A locator map indicates the position of the Coquimbo Region within Chile and South
 786 America.

787

788

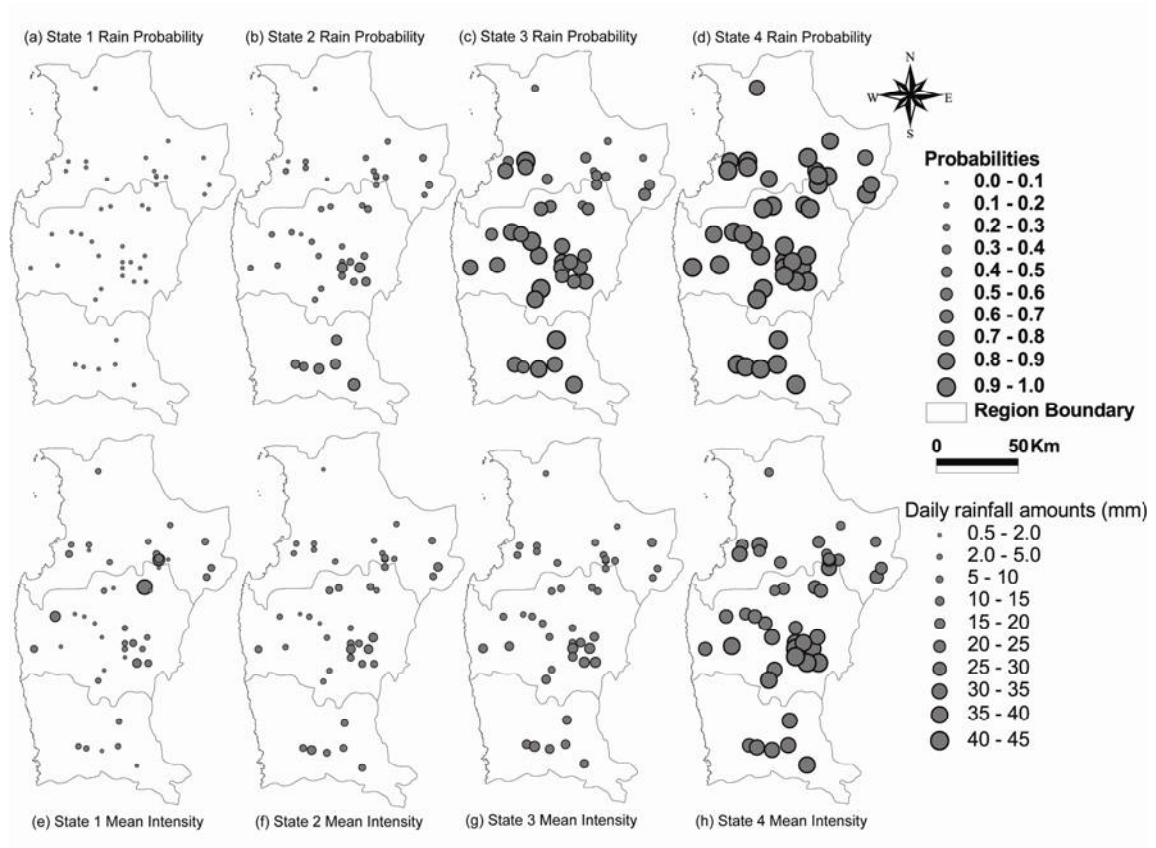
789



790

791 FIG. 4. Spatial correlation function of a) daily rainfall amount, b) rainfall intensity and c)
 792 rainfall frequency for the daily, 2-day, 5-day, 10-day, 30 day and seasonal averaging periods.
 793 The 1/e decorrelation value is shown as a dashed line.

794

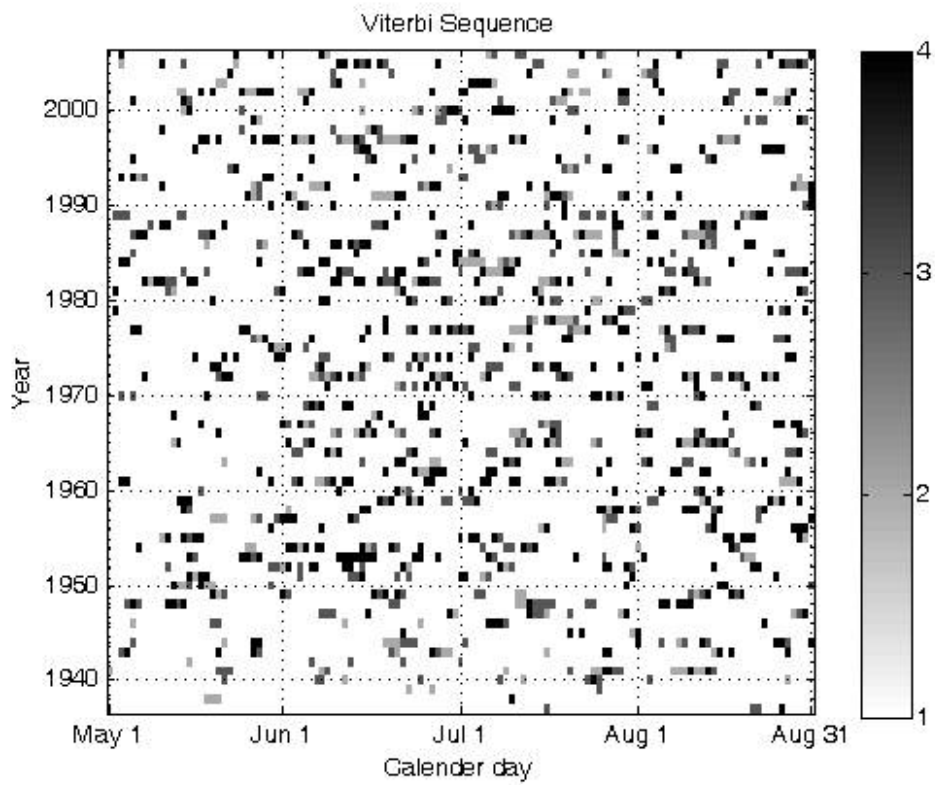


795

796 FIG. 5. Four-state HMM rainfall parameters. (a)–(d): probabilities of rainfall occurrence and

797 (e)–(h): mean rainfall intensities (i.e. wet-day amounts).

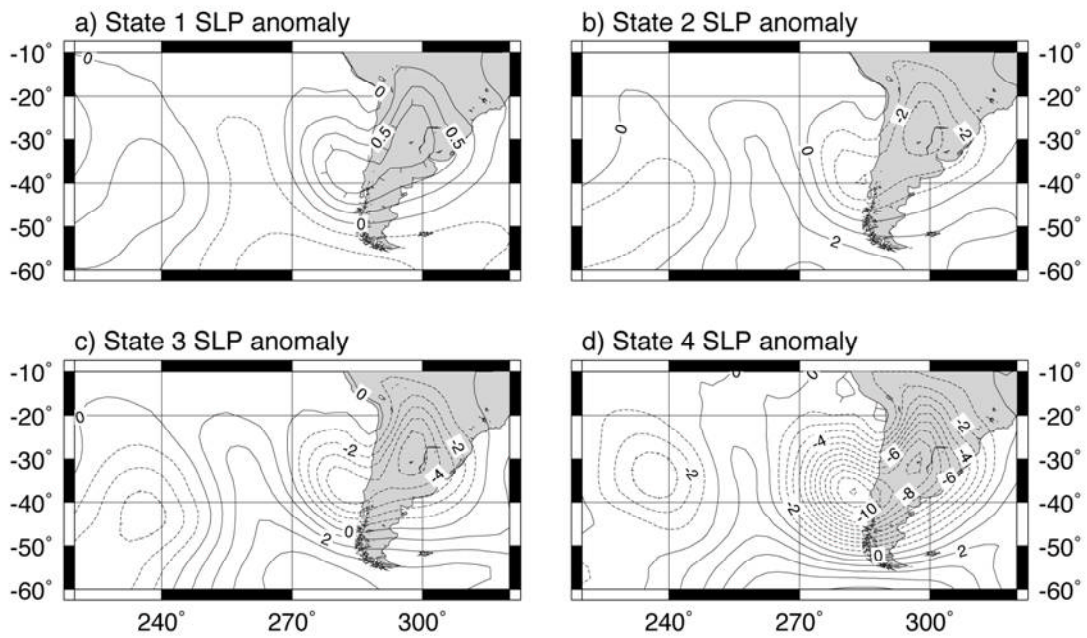
798



799

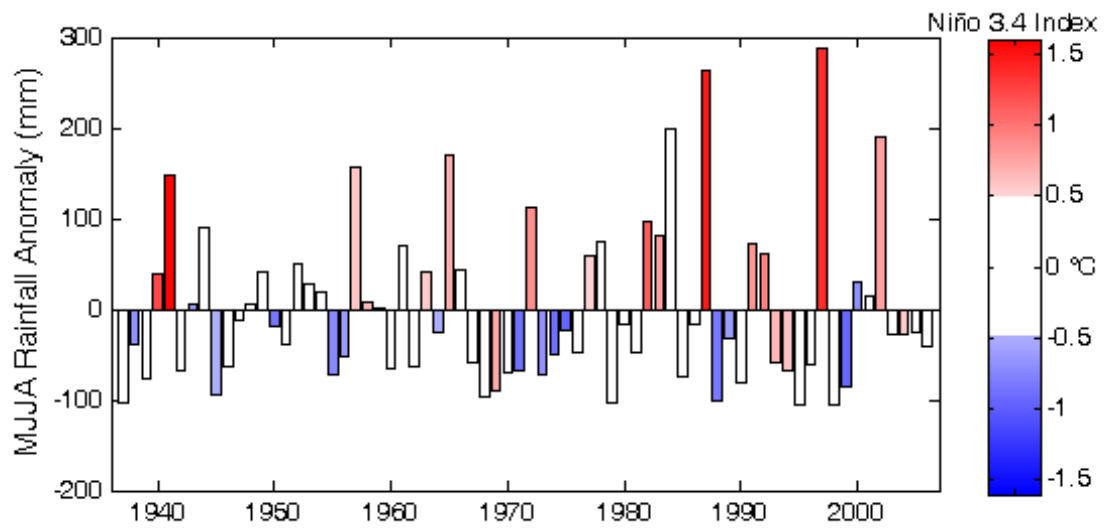
800 FIG. 6. The most probable HMM state sequence obtained using the Viterbi algorithm. Rainfall
 801 states are indicated from driest (state 1) to wettest (state 4) on the grey color bar.

802



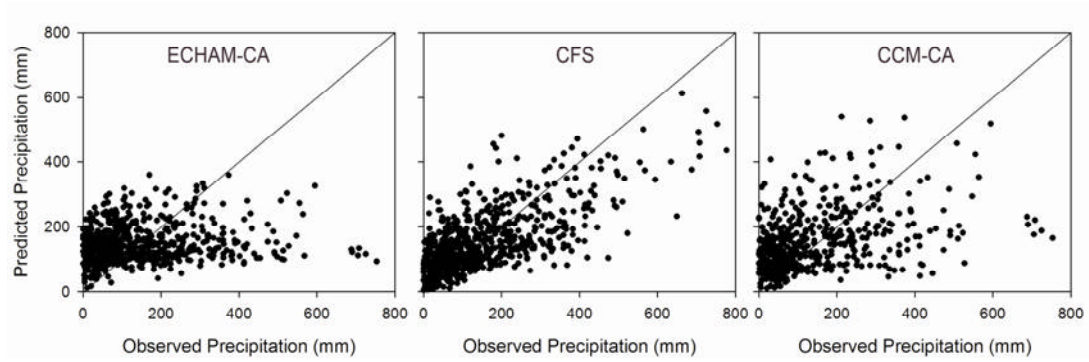
803

804 FIG. 7. Composites of sea-level pressure anomalies (mb) for each rainfall state. A finer
 805 contour interval was used in panel (a) for clarity.



806

807 FIG. 8. Station-average MJJA rainfall amount, colored according to the sign and magnitude of
 808 the Niño 3.4 SST index for the period 1937–2005.

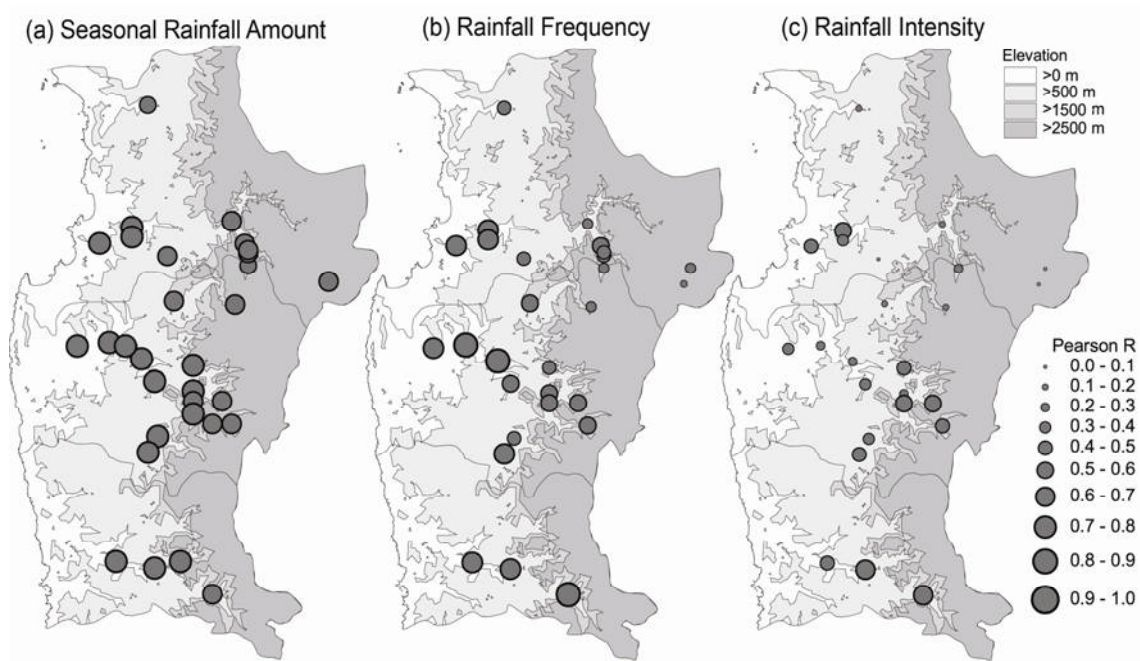


809

810 FIG. 9. Cross-validated hindcasts versus observed precipitation amounts using CCA for the
811 three GCMs for the period 1981–2000, where each circle represents the value for each
812 station, for each year. Thus there are 42x20 circles in each panel.

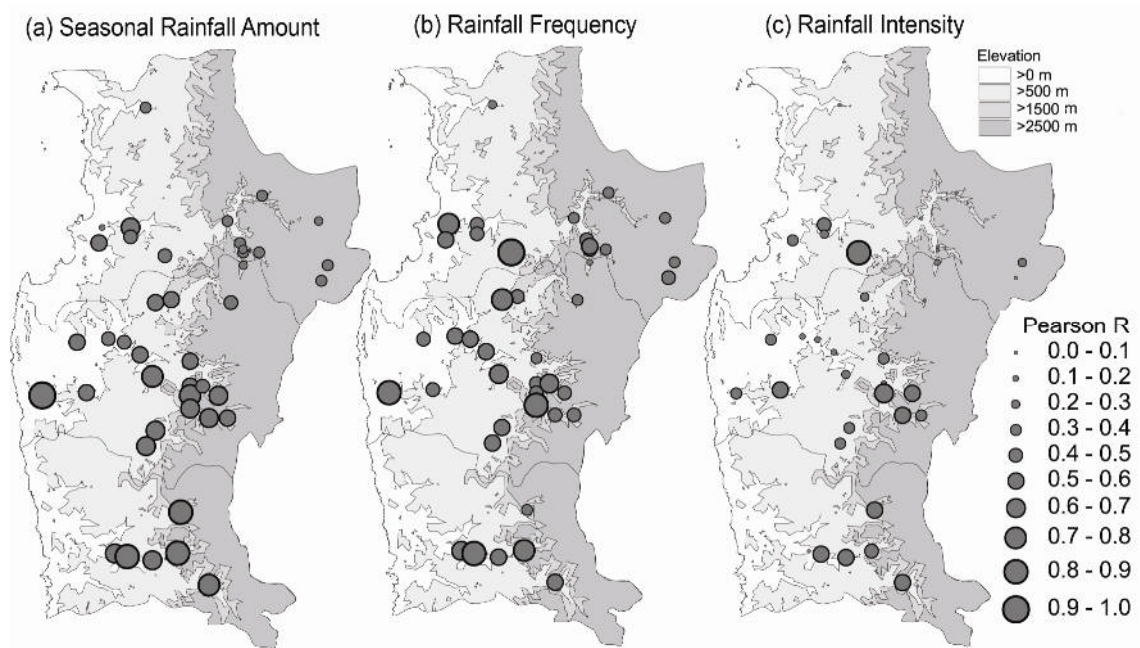
813

814



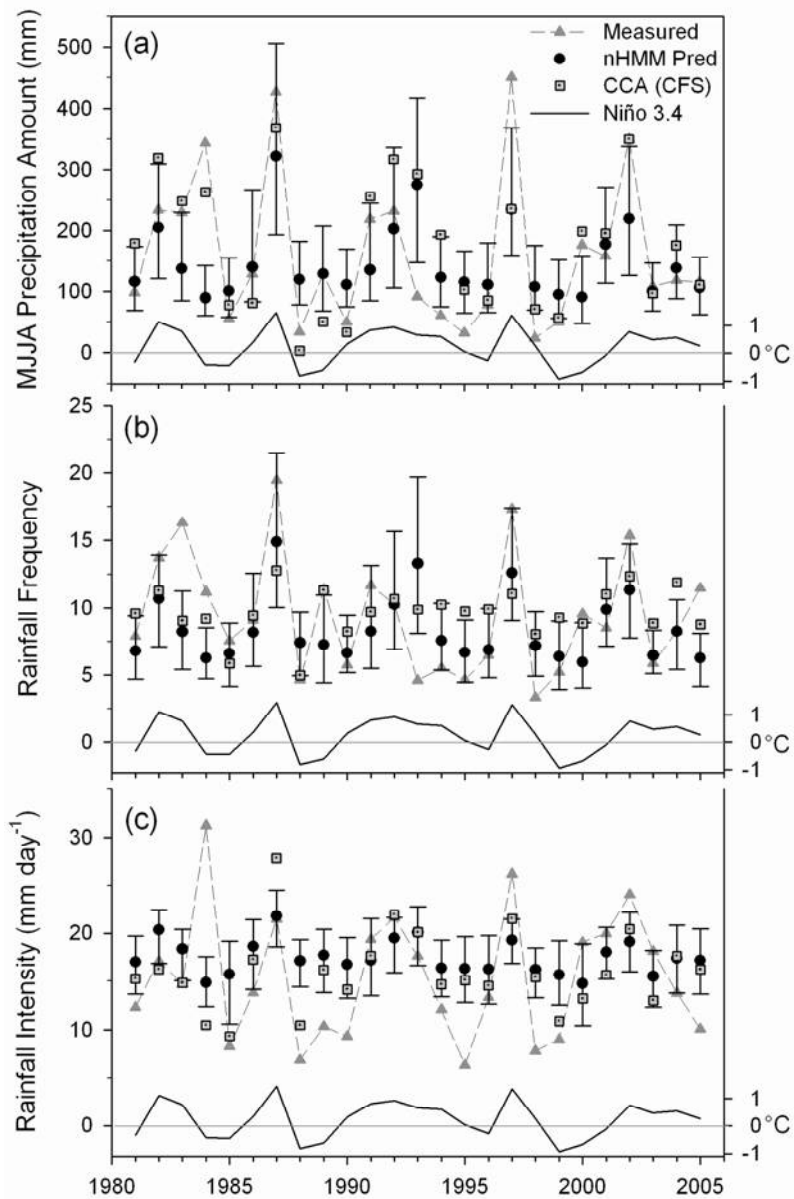
815

816 FIG. 10. Pearson correlation between CFS hindcasts downscaled using CCA and observed
817 rainfall for (a) seasonal rainfall amount, (b) rainfall frequency, and (c) mean rainfall intensity,
818 for the period 1981–2005.



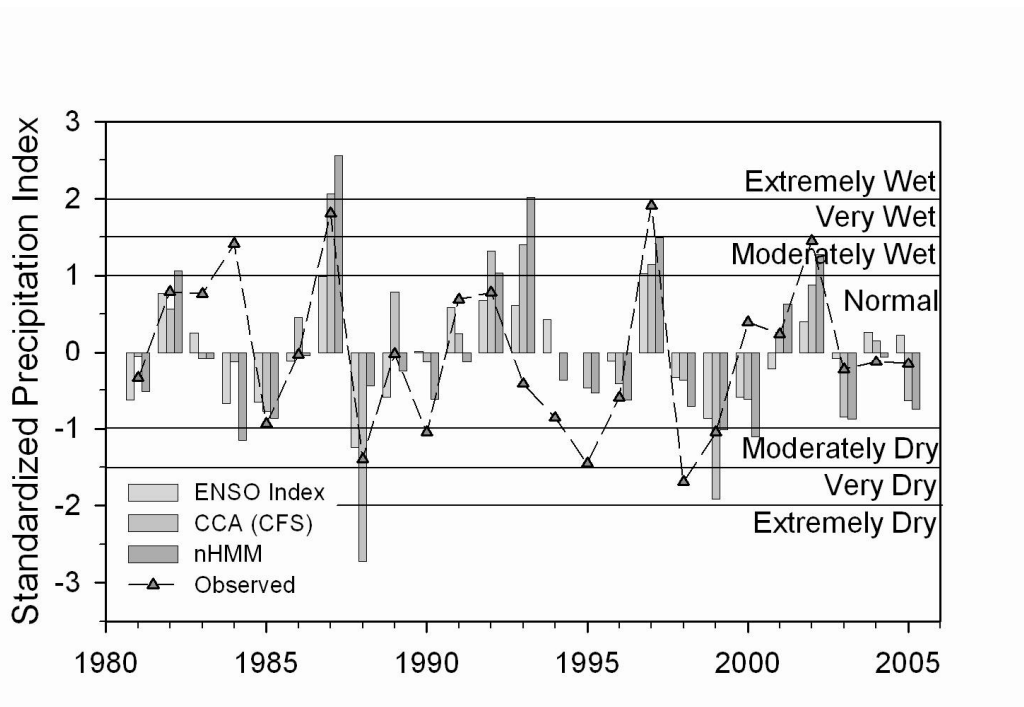
819

820 FIG. 11. As Fig. 10, but for CFS downscaled rainfall obtained using the nHMM and taking the
 821 ensemble mean over the 150 nHMM simulations.



822

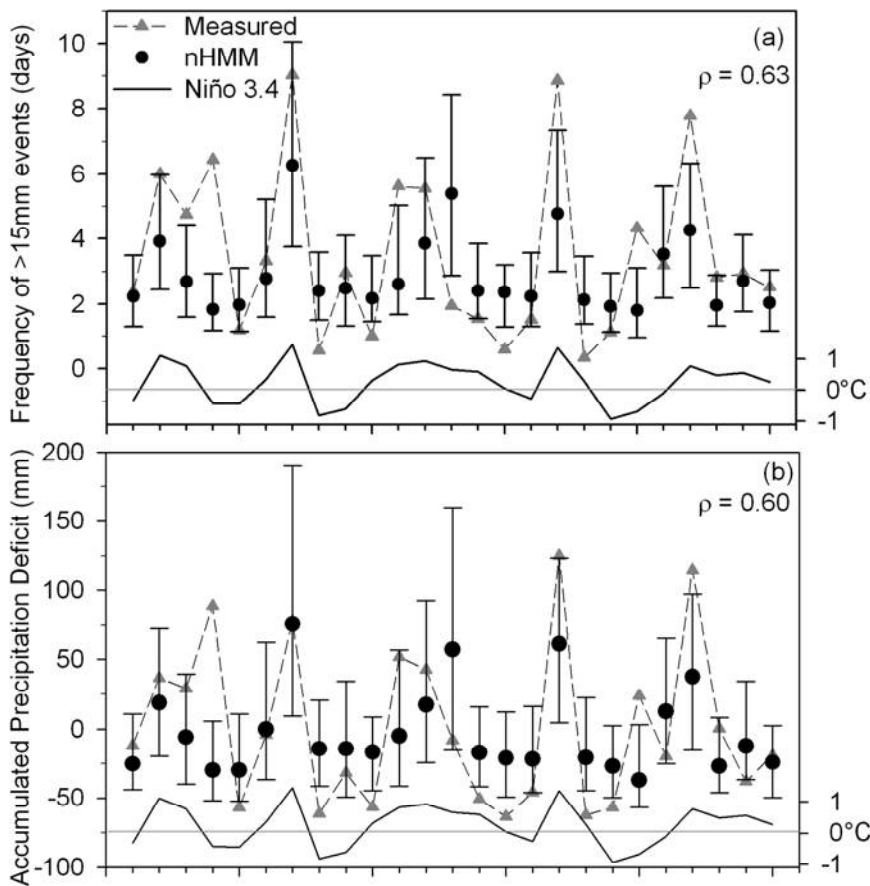
823 FIG. 12. Comparison of station-averaged downscaling results obtained from CFS using CCA
 824 and the nHMM. (a) seasonal rainfall amount, (b) rainfall frequency (c), mean rainfall
 825 intensity. The error bars indicate the 25th and 75th percentiles of the simulated nHMM
 826 ensemble values. The Niño 3.4 index is also indicated.



828

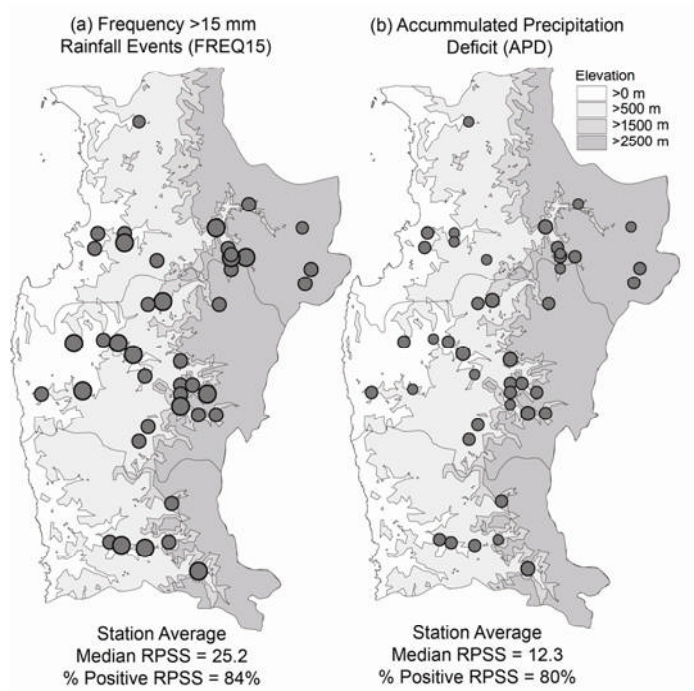
829 FIG. 13. Station-averaged hindcasts of the Standardized Precipitation Index (SPI) obtained
 830 from the MJJA Niño 3.4 Index, and the CFS downscaled with CCA and with the nHMM. The
 831 SPI values constructed from observed station rainfall are also plotted.

832



833

834 FIG. 14. Hindcasts of two station-averaged meteorological drought indices (circles) consisting
 835 of (a) the number of days with rainfall exceeding 15mm (FREQ15), and (b) the Accumulated
 836 (daily) Precipitation Deficit (APD), based on the nHMM simulations, compared with values
 837 constructed from observed daily rainfall (dashed). The error bars indicate the 25th and 75th
 838 percentiles of the nHMM ensemble. The Niño 3.4 index is also indicated.



839

840 FIG. 15. Median Rank Probability Skill Score (RPSS) for hindcasts of a) the number of days
 841 with rainfall exceeding 15mm (FREQ15) and b) the Accumulated Precipitation Deficit
 842 (APD), constructed from the nHMM simulations.